#### Least Angle Regression (LARS)

#### LASSO

- LASSO is a constrained version of Ordinary Least Squares (OLS)
- Let  $x_1, x_2, \dots x_p$  be the variables/predictors and y be the response
- $X = \begin{bmatrix} x_1 & x_2 \dots x_p \end{bmatrix}$ , the matrix with columns containing the predictors.
- If the regression coefficients,  $\hat{\beta} = (\hat{\beta}_1, \hat{\beta}_2 \dots \hat{\beta}_p)'$  give the estimated response  $\hat{y}$ , then

$$\hat{y} = \sum_{j=1}^{p} x_j \hat{\beta}_j$$

• LASSO chooses  $\hat{\beta}$  by minimizing total squared error subject to a constraint on the coefficients, i.e

$$\min \left| |y - \hat{y}| \right|^2 \quad subject \ to \ \sum_{j=1}^p |\hat{\beta}_j| \le t$$

## Forward Stagewise Regression

- A subset selection method (select a subset of variables with linear regression)
- Idea:
- Repeat
  - Select the predictor having largest absolute correlation with residual vector
  - Update the estimated response to move in the direction of the correlated variables

## Stagewise

- Iterative technique that begins with  $\hat{y} = 0$  and builds up the regression function in successive small steps.
- At any step, let  $c(\hat{y})$  be the vector of *current correlations*  $\hat{c} = c(\hat{y}) = X'(y - \hat{y})$

where  $\hat{c}_j$  is proportional to the correlation between predictor  $x_j$  and current residual vector  $(y - \hat{y})$ 

• The Stagewise algorithm moves the prediction in the direction of the greatest current correlation

$$\hat{j} = argmax[\hat{c}_j] \text{ and } \hat{y} = \hat{y} + \epsilon * sign(\hat{c}_j) * x_j$$

where  $\epsilon$  is a small constant

### Least Angle Regression

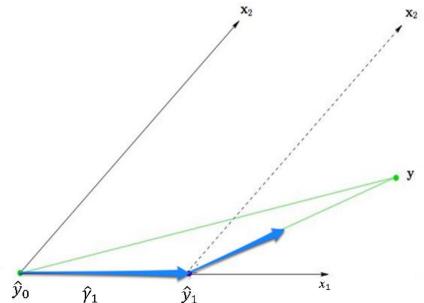
- Both LASSO and Stagewise are variants of a basic procedure called Least Angle Regression (LARS)
- LARS is a stylized version of the Stagewise procedure
- Only *p* steps are required for the full set of solutions (*p* is the number of variables) for LARS
- The advantages of LARS compared to Stagewise procedure
  - The prediction movement is in a direction which is equiangular to all the most equally correlated variables (**optimal directions**)
  - optimum stepsize (i.e  $\epsilon$ ) is taken such that the next variable is equally correlated with the previously taken variables (**optimal sized leaps**)

# Least Angle Regression

- Assume standardized predictors in the model (mean 0 and unit variance)
- Algorithm
  - Start with no predictors in the model
  - Find the predictor  $x_1$  most correlated to the residual (equivalently, the variable making **least angle** with the residual)
  - Keep moving in the direction of the most correlated predictor until another predictor  $x_2$  becomes equally correlated with the residual.
  - Move in a direction equiangular to both the predictors
  - Continue until all the predictors are in the model

# Example

- Example with 2 correlated variables  $x_1$  and  $x_2$
- y is the response,  $\hat{y}_0$ ..  $\hat{y}_2$  are estimated responses at each step ,  $\hat{\gamma}_1 \dots \hat{\gamma}_2$  are the optimal step sizes
- At  $\hat{y}_0 = 0$ , the residual vector  $y \hat{y}_0$  is most correlated to  $x_1$  (least angle)  $\hat{y}_1 = \hat{y}_0 + \hat{\gamma}_1 x_1$
- Select  $\hat{\gamma}_1$  such that the residual  $y-\hat{y}_1$  is equally correlated with  $x_1$  and  $x_2$
- Then,  $\hat{y}_2 = \hat{y}_1 + \hat{\gamma}_2 u_2$ , where  $u_2$  is the unit bisector (equiangular vector)
- Here,  $\hat{y}_2 = y$



## LARS (contd)

- At  $k^{th}$  step, let  $A_k$  be the set of active variables/predictors and  $\beta_{A_k}$  be the coefficient vectors and  $\hat{y}_k$  be the estimated response.
- $X_{A_k}$  be the active variables i.e those variables whose absolute correlation with the residual vector equals the maximal achievable absolute correlation.
- The current residual will be  $r_k = y \hat{y}_k = y X_{A_k}\beta_{A_k}$
- The coefficients are moved in the direction  $\delta_k = (X_{A_k}^T X_{A_k})^{-1} X_{A_k}^T r_k$

i.e 
$$\beta_{A_k}(\gamma) = \beta_{A_k} + \gamma \delta_k$$

#### Equiangular vector

• 
$$\beta_{A_k}(\gamma) = \beta_{A_k} + \gamma \delta_k$$

• Thus 
$$\hat{y}(\gamma) = X_{A_k}\beta_{A_k}(\gamma) = X_{A_k}\beta_{A_k} + \gamma * X_{A_k}\delta_k$$

 $u_k$ : makes equal angles with the predictors in  $A_k$ 

- Prove that  $u_k$  makes equal angles with the active predictors  $X_{A_k}^T u_k = X_{A_k}^T X_{A_k} \delta_k = X_{A_k}^T X_{A_k} (X_{A_k}^T X_{A_k})^{-1} X_{A_k}^T r_k = X_{A_k}^T r_k$
- The entries of the vector  $X_{A_k}^T r_k$  are all the same since the predictors have same correlation with residuals (by construction)

## LAR algorithm:

Least Angle Regression is similar to forward stagewise, but only enters "as much" of a predictor as it deserves

Algorithm 3.2 Least Angle Regression.

- 1. Standardize the predictors to have mean zero and unit norm. Start with the residual  $\mathbf{r} = \mathbf{y} \bar{\mathbf{y}}, \beta_1, \beta_2, \dots, \beta_p = 0.$
- 2. Find the predictor  $\mathbf{x}_j$  most correlated with  $\mathbf{r}$ .
- 3. Move  $\beta_j$  from 0 towards its least-squares coefficient  $\langle \mathbf{x}_j, \mathbf{r} \rangle$ , until some other competitor  $\mathbf{x}_k$  has as much correlation with the current residual as does  $\mathbf{x}_j$ .
- 4. Move  $\beta_j$  and  $\beta_k$  in the direction defined by their joint least squares coefficient of the current residual on  $(\mathbf{x}_j, \mathbf{x}_k)$ , until some other competitor  $\mathbf{x}_l$  has as much correlation with the current residual.
- 5. Continue in this way until all p predictors have been entered. After  $\min(N-1, p)$  steps, we arrive at the full least-squares solution.

# Modification of LAR to solve LASSO problem

Algorithm 3.2a Least Angle Regression: Lasso Modification.

4a. If a non-zero coefficient hits zero, drop its variable from the active set of variables and recompute the current joint least squares direction.

## References

- Efron, Bradley, et al. "Least angle regression." *The Annals of statistics* 32.2 (2004): 407-499.
- The Elements of Statistical Learning: Data Mining, Inference, and Prediction, Trevor Hastie, Robert Tibshirani, Jerome Friedman, Springer, 2008